

1 **A Joint Econometric Model Framework for Transportation Network Companies (TNC)**
2 **Users' Trip Fare and Destination Choice Analysis**

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1 ABSTRACT

2 In this study, we examine the factors affecting Chicago Transportation Networking Companies
3 (TNC) pricing and destination choice behavior. While trip fare has been examined from various
4 perspectives, earlier fare models have not considered an exhaustive set of independent variables.
5 Further, trip fare decisions are significantly influenced by trip destination. Hence, in our study a
6 joint model of trip fare and destination choice is proposed. The joint model system – linear
7 regression for fare and multinomial logit model for destination - is developed based on Chicago
8 TNC weekday trip data from January 2019 to December 2019. A wide range of origin and
9 destination specific land use and built environment factors, transportation infrastructure attributes,
10 and weather attributes were found to be significant in the model system. Based on log-likelihood
11 (LL) and Bayesian Information Criterion (BIC) measures, the model performance of the proposed
12 joint model is found to be superior compared to independent fare and destination models. The
13 applicability of our proposed fare and destination choice model is illustrated through fare
14 prediction and destination elasticity analysis. The framework can potentially be employed to
15 generate TNC fare for inclusion in Level of Service measures for TNC model in the mode choice
16 model.

17
18 **Keywords:** Transportation Networking Companies (TNC), Joint Linear Regression (LR) and
19 Multinomial Logit (MNL) Model, Prediction, Elasticity Analysis

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1 INTRODUCTION

2 Transportation Networking Companies (TNCs) are reshaping the transportation sector with
3 operations in more than 10,000 cities across the world (1). As of 2021, the global ride share market
4 is valued at 85.8 billion and is predicted to be valued at 185.1 billion by the end of 2026 (2). In a
5 recent report, TNC heavyweight Uber (3) indicated that about 118 million users have used Uber
6 service at least once a month in 2021. The magnitude of the user base, considering the ongoing
7 COVID-19 pandemic in 2021, is illustrative of the major influence of TNC on mobility in urban
8 regions. TNCs are emerging as makeshift public transport options across many urban regions
9 across the world including Chicago (4), San Francisco (5), Boston (6), Santiago, Chile (7),
10 Chengdu, China (8), and Hanover, Germany (9). As TNCs become an increasingly significant
11 transportation mobility alternative across the world, there is growing literature examining TNC
12 impact on the various facets of the transportation system.

13 An important consideration with the growing adoption of TNC alternatives is the inclusion
14 of these systems within urban travel demand modeling frameworks. Several research efforts have
15 examined the impact of TNCs in the context of demand generation and distribution across the
16 urban region. However, incorporating TNCs within the current mode choice frameworks across
17 urban regions is not typically explored. The main reason it is challenging to develop mode choice
18 frameworks for TNCs is the lack of an easy to adopt framework for generating Level of Service
19 (LOS) measures. The generation of travel time measure is relatively easy as automobile travel
20 times can be directly applied for TNC travel times (for solo passengers). However, generating the
21 cost measure is not straightforward. The proposed study is geared towards tackling this challenge
22 of predicting trip level TNC fare that can be incorporated within travel demand model frameworks
23 for generating travel cost measure for TNC alternatives in mode choice. In modeling trip fare, the
24 current study postulates that TNC user's selection of a trip is closely linked with destination and
25 the associated fare. For example, the destination attractiveness of a location with high density of
26 hospitality venues (hotels/motels) is quite high. At the same time, the fare to such destinations
27 might also be higher due to the demand. This is an example of how a destination attribute affects
28 fare and destination choice. These can be readily considered in fare and destination models.
29 However, it is also likely that factors such as local events (such as a concert) occurring in a
30 destination might affect fare and demand. The information on such events might not be available
31 for modeling. Hence, the influence of such unobserved information can be considered in the form
32 of common unobserved factors affecting fare and destination. Further, given the ease with which
33 TNC rides can be selected on smartphone apps, it is possible that TNC users can revisit their choice
34 of destination in response to the fare levels shown in the app. With these considerations, in our
35 study, we develop a joint model of fare and destination choice where trip fare is modelled using
36 linear regression model (LR) and destination choice is modelled using a multinomial logit model
37 (MNL). The model estimation exercise is conducted using TNC data from Chicago region.
38 Specifically, weekday trip data spanning January 2019 through December 2019 is employed for
39 our analysis. Trip fare and destination data are further augmented with a host of independent
40 variables including trip attributes, origin attributes, destination attributes, land use and built
41 environment attributes, socio-demographic attributes, and weather attributes. The model
42 estimation process is augmented by elasticity analysis to illustrate how the proposed model can be
43 employed to understand the influence of various independent variables on fare and destination
44 selection.

45 The rest of the paper is organized as follows: Literature review section summarizes relevant
46 literature and positions the current study. Data section documents the data processing procedures

1 and provides an overview of the data used in our analysis. The mathematical details of the models
2 are described in the following section. Model Estimation Results section describes the results from
3 the models. An elasticity analysis illustrating the impact of independent variables is documented
4 in the next section. Conclusions section presents an overview of the paper and identifies potential
5 directions for future research.

7 **LITERATURE REVIEW AND CURRENT STUDY IN CONTEXT**

8 We present an overview of earlier research efforts on the two TNC dimensions of interest in our
9 research – trip fare and destination.

10 TNC fare is evaluated in two ways in earlier research. *First*, fare is considered as an
11 independent variable affecting the decision to use TNC alternatives. In these studies, various TNC
12 associated decisions such as solo or pooled trip (10, 11), competition between transit and TNC (5,
13 12, 13), role of income in affecting TNC usage (12, 14, 15), driver economics and turnover (16–
14 18) and satisfaction with TNC (12, 19, 20) are examined. Important findings from these studies
15 include: (a) high income individuals prefer TNC to transit (12, 15), (b) higher TNC pricing power
16 is observed in highly walkable areas (21, 22), (c) turnover for ridehailing services is significantly
17 high (16), (d) sharing TNC demand and supply information with drivers may lead to higher
18 satisfaction level among drivers (20), and (e) a higher inclination among younger individuals for
19 using TNC (5, 12, 23). *Second*, studies examined dynamic pricing policy (or surge pricing) in their
20 analysis of TNC systems. In these studies, fare is modeled as a continuous variable within an
21 optimization framework (24–26). The approaches provide elegant mathematical formulations for
22 profit maximization or demand imbalance minimization in the context of a equilibrium based
23 optimization models to estimate price and/or demand. The mathematical formulations are
24 applicable under a host of assumptions such as restricted number of TNCs (25), neglecting spatial
25 variations (27), the distances in the network are equidistant (24), and limits on the number of modal
26 alternatives (for example only Drive vs TNC in Afifah and Guo (25)). The demand, price and
27 model choice equations in these approaches are simplified and focus on a small set of variables
28 such as trip length (11). While these approaches are very helpful, applying these methods for large
29 urban regions with temporal and spatial variations are not readily practical. In our review, we
30 found only 3 studies that developed direct fare models using TNC data (11, 28, 29) where a small
31 set of variables such as trip distance, trip time, tolls and additional charges were considered.

32 Destination selection behavior has been examined in multiple ride sharing domains
33 including bicycle-sharing system (30–32), taxi (33–35), TNC and Shared Autonomous Vehicle
34 (SAV) (36, 37). The preferred approach employed at the disaggregate level is the Multinomial
35 Logit Model (MNL) based on the random utility maximization approach (30). Other model
36 structures employed for analysis of destination dimensions such as Traffic Analysis Zone (TAZ)
37 (38) includes a Generalized Spatially Correlated Logit (GSCL) Model. In some studies, aggregate
38 destination allocations are analyzed using Multiple Extreme Continuous Extreme Value
39 (MDCEV) models (32). Important findings on destination choice preferences include: (a)
40 destination choice is highly correlated with employment status (39), (b) presence of high demand
41 in the neighborhood is a strong contributor of demand (32), (c) lower fare price increases the utility
42 of a destination (40), (d) duration of stay and home location prior to the activity affect destination
43 choice (41), and (e) destination choice behavior is influenced by the perceived destination image
44 from individual's social network (42).

45

1 **Contributions of the Current Study**

2 Several studies have recognized that pricing algorithms are influenced by spatio-temporal demand
3 (such as demand at origin in preceding 15 minutes), origin and destination land use and built
4 environment factors, transportation infrastructure attributes, and weather attributes (24, 43).
5 However, none of the earlier research studies have incorporated a wide range of attributes in
6 modeling TNC fare. The first contribution of our study is to develop a comprehensive trip fare
7 model while accounting for a host of independent variables. In this study, we recognize that trip
8 fare values are closely aligned with trip destination. Hence, the second contribution of our study
9 is to develop a joint model system that accounts for common unobserved factors affecting fare and
10 destination. The study develops a joint linear regression (LR) for fare and multinomial logit (MNL)
11 model for destination labelled as the LR-MNL model. The model system is developed using TNC
12 trip data from Chicago for the year 2019. Chicago data has been employed in the literature to
13 study various TNC dimensions including spatial demand variations and willingness to use pool
14 alternative (10, 21, 44). Finally, the current study contributes empirically by allowing us to
15 understand Chicago TNC pricing model and destination choice behavior. The framework can
16 potentially allow us to generate TNC fare for mode choice model. In application, the model
17 developed can be employed in a sequence – destination choice outcome followed by trip fare
18 prediction. The model framework can also allow us to identify systemic differences across the
19 Chicago city in pricing (if any) and how various destination attributes influence destination
20 preferences.

21

22 **DATA PREPARATION**

23 **Data Source**

24 City of Chicago has made TNC data available for analysis beginning in November 2018. As of
25 2019, three TNCs were operating in the Chicago area: Uber, Lyft and Via (45). For this current
26 study, daily weekday trip data of more than 50 million records for 12 months starting from January
27 2019 to December 2019 was compiled for our analysis(45). Origin and destination for each of
28 these trips have been aggregated at the census tract level while trip times (start time & end time),
29 trip fare are rounded to nearest 15 minutes and 2.50 USD respectively. The trip dataset is further
30 augmented by trip attributes such as trip start & end time, trip distance, shared trip indicator
31 provided by Transportation Network Providers-Chicago Data Portal (45), land use and built
32 environment variables including distance from Central Business District (CBD), residential area,
33 commercial area, institutional area, recreational area accessed from Chicago Data portal and
34 Chicago Metropolitan Agency for Planning (CMAP) (44, 46), Transportation infrastructure
35 attributes including bike lane density, street length, number of bus stops, number of transit stations,
36 number of divvy stations walk score, transit score compiled from Chicago Data portal and Chicago
37 Metropolitan Agency for Planning(CMAP) (45, 46) and sociodemographic attributes such as low
38 income indicator, employment density drawn from US Census Bureau (47) and weather attributes
39 such as snow depth obtained from National Climatic Data Center (NCDC) (48). A summary of the
40 independent variables is provided in Table 1.

41

42 **Sample Formation**

43 The data processing procedures were implemented in the following sequence. First, records with
44 missing and inconsistent information were dropped from the dataset. Second, trips that originated
45 or destined outside of Chicago city area were removed from the dataset. Finally, weekday trips
46 were retained amounting to more than 44 million of records. The spatial distribution of weekday

trips by origin and destination census tract are presented in **Figure 1(a)** and **Figure 1(b)** respectively. Employing the full set of records (44 million) would increase computational time for modeling exercise significantly. Further, using such large datasets in econometric models might lead to overfitting. To address these issues, we randomly select 25 samples of 10,000 records for our model estimation exercise. These samples will allow us to ensure that the parameters estimated using one sample are not significantly different from other samples of data. Towards this end, we conduct a rigorous statistically valid comparison of model estimates across all 25 samples prior to selecting a sample for further analysis.

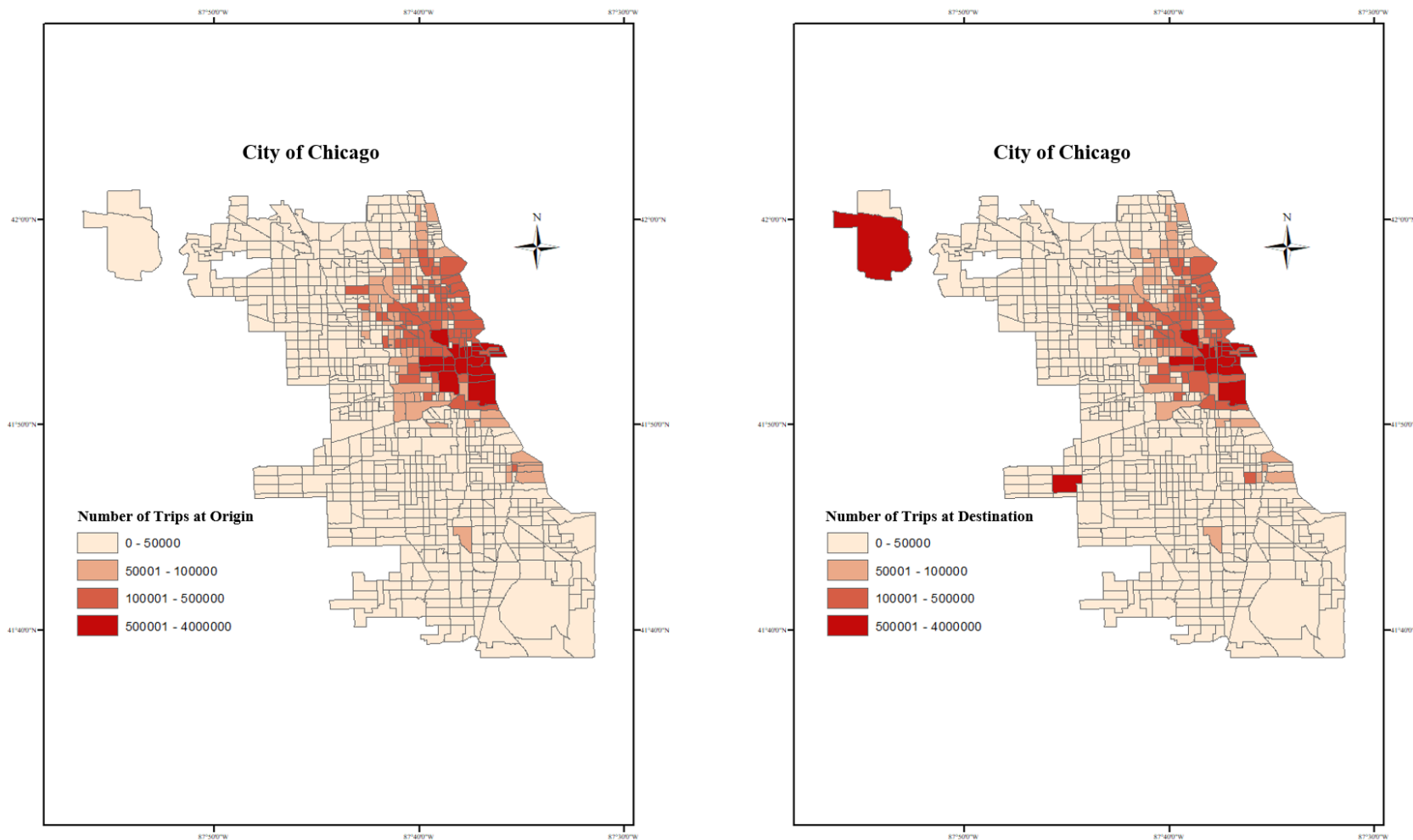
For the destination choice models, all census tracts in the region are potential alternatives. In our data for Chicago we identified 801 census tracts (49). From this broad set of alternatives, destination choice models are developed employing a random sample of 30 alternatives (inclusive of the chosen alternative). Similar random sampling process has been adopted in earlier literature for destination choice models(see 57–60 for details).

TABLE 1 Descriptive Statistics of Variables

| Variables | Variable Descriptions | Descriptive Statistics | |
|--|--|------------------------|-----------|
| | | Mean | Std. dev. |
| DEPENDENT VARIABLES | | | |
| Trip fare model | | | |
| Trip fare | Ln (Trip fare) | 2.079 | 0.577 |
| INDEPENDENT VARIABLES (CONTINUOUS) | | | |
| Trip Attributes | | | |
| Trip distance | Distance traveled in each trip | 4.149 | 4.129 |
| Network distance | Ln (Shortest distance between census tracts) | 1.895 | 0.494 |
| Demand in last 15 minutes at origin | Ln (Demand in last 15 minutes in each origin census tract) | 2.006 | 1.482 |
| Demand in last 15 minutes at destination | Ln (Demand in last 15 minutes in each destination census tract) | 2.032 | 1.537 |
| Land Use and Built Environment Attributes | | | |
| Network distance from CBD | Ln (Network distance to census tract from Central Business District (CBD)) | 1.871 | 0.560 |
| Residential area | Total residential area in each census tract (area/100) in acre | 0.602 | 0.519 |
| Commercial area | Total commercial area in each census tract (area/100) in acre | 0.115 | 0.174 |
| Institutional area | Total institutional area in each census tract (area/100) in acre | 0.113 | 0.280 |
| Recreational area | Total recreational area in each census tract (area/100) in acre | 0.074 | 0.232 |
| Land use mix | Land use mix = $\frac{-\sum_k(P_k(\ln P_k))}{\ln N}$, where k is the category of land-use, p is the proportion of the developed land area for specific land-use, N is the number of land-use categories | 0.134 | 0.045 |
| Transportation Infrastructure Attributes | | | |

| Variables | Variable Descriptions | Descriptive Statistics | |
|--|--|------------------------|------------|
| | | Mean | Std. dev. |
| Bike lane density | Length of bike lane in each census tract per acre (Density*100) (mi/acre) | 0.321 | 0.352 |
| Length of street | Length of street in each census tract | 5.597 | 4.953 |
| Number of bus stops | Number of bus stops in each census tract | 12.486 | 8.330 |
| Number of L stations | Number of stations of L transit system in each census tract | 0.156 | 0.529 |
| Number of divvy stations | Number of divvy stations in each census tract | 1.029 | 1.441 |
| Walk score | Walk score (a measure of serviceability of walkability) in each census tract | 82.397 | 26.225 |
| Transit score | Transit score (a measure of serviceability of public transit) in each census tract | 8.260 | 0.996 |
| Sociodemographic Attributes | | | |
| Employment density | Number of employments in each census tract per acre (Density/100) | 0.236 | 0.381 |
| Weather Attributes | | | |
| Snow depth | Standard score ($\frac{x-\mu}{\sigma}$) of snow depth in each census tract. Where x is the observed value of snow depth, μ is the mean of the distribution of the values of snow depth and σ is the standard deviation of the distribution of the values of snow depth | 0.004 | 1.034 |
| INDEPENDENT VARIABLES (CATEGORICAL) | | | |
| Variables | Variable Descriptions | Freq. | Percentage |
| Trip Attributes | | | |
| Trip starts at AM peak | Trip starts within AM peak period | 1965.000 | 19.650 |
| Trip starts at PM peak | Trip starts within PM peak period | 2560.000 | 25.600 |
| Trip starts at other time | Trip starts in other time period | 5475.000 | 54.750 |
| Trip ends at AM peak | Trip ends within AM peak period | 1876.000 | 18.760 |
| Trip ends at PM peak | Trip ends within PM peak period | 2481.000 | 24.810 |
| Trip ends at other time | Trip ends in other time period | 5643.000 | 56.430 |
| Shared trip indicator | | | |
| Yes | Trip authorized as shared | 1507.000 | 15.070 |
| No | Trip is not authorized as shared | 8493.000 | 84.930 |
| Sociodemographic Attributes | | | |
| Low income indicator | | | |
| Yes | Census tract with median income under \$58 thousand USD (15th percentile) | 466.000 | 58.543 |
| No | Census tract with median income over \$58 thousand USD (15th percentile) | 330.000 | 41.457 |

1



2
3 **Figure 1 Total number of weekday trips (a) originated; (b) destined**

1 **ECONOMETRIC METHODOLOGY**

2 In this study, we develop a joint trip fare and trip destination model where trip fare is modelled
 3 using a linear regression model and trip destination is modelled using a multinomial logit model.
 4 Let, q ($=1, 2, 3, \dots, Q=10,000$) be an index to represent each individual trip, y_q be an index to
 5 represent the fare associated with a trip q , and s ($=1, 2, \dots, S=30$) be an index to represent
 6 destination alternatives (census tracts). In the following sections, we describe two model
 7 components and then present estimation procedure for the joint model.

8 **Trip Fare Model**

9 In the linear regression formulation, we express y_q as a function of independent variables z_q as
 10 follows:

$$11 \quad y_q = (\alpha' + \eta')z_q + \varepsilon_q \quad (1)$$

12 where α' is a vector of coefficients to be estimated, η represents the effect of common
 13 unobserved factors modifying the impact of z_q in the trip fare and trip destination models (see
 14 Equation 2) and ε_q is an idiosyncratic random error term assumed independently normally
 15 distributed with variance γ^2 . Now, we can express the probability of a trip, q having fare, y_q as
 16 follows:

$$17 \quad P(y_q) = \frac{\phi \left[\frac{y_q - (\alpha' + \eta')z_q}{\gamma} \right]}{\gamma} \quad (2)$$

18 where $\phi(\cdot)$ is the standard normal probability distribution function.

19 **Trip Destination Model**

20 In the MNL model, the random utility of an alternative s for trip q takes the following form:

$$21 \quad u_{qs} = (\beta' + \eta')x_{qs} + \epsilon_{qs} \quad (3)$$

22 where u_{qs} is the utility obtained by user q by choosing census tract s as the destination
 23 from a choice set of 30 census tracts. x_{qs} is a vector of attributes and β is a vector of model
 24 coefficients to be estimated. The random error term, ϵ_{qs} , is assumed to be independent and
 25 Gumbel-distributed identically across the dataset. In random utility maximization (RUM)
 26 approach, a user making the trip, q will choose a census tract as the destination that offers the
 27 highest utility. Therefore, the probability expression takes the following multinomial logit form:

$$28 \quad P(s_q) = \frac{\exp((\beta' + \eta')x_{qs})}{\sum_{s=1}^S \exp((\beta' + \eta')x_{qs})} \quad (4)$$

29 The destination alternatives in our study context are not labelled (i.e., they are not typical
 30 categorical alternatives such as travel mode (car, bike)). Hence, our model estimation approach
 31 considers a generic parameter structure across all alternatives. The approach will allow for
 32 parameter estimation for variables that vary across destination alternatives such as destination
 33 employment or destination land use mix. In the model structure, accounting for variables at the
 34 trip level such as trip start time or origin destination can be considered as an interaction term with
 35

1 variables varying across the destination (such as Trip starts in AM peak x Number of divvy stations
2 in CT).

3

4 **Estimation Procedure**

5 To complete the model structure of the Equations (1) and (3), it is necessary to define the structure
6 for the unobserved vector η . In this paper, we assume that this vector is independent realizations
7 from normal distributions as follows: $\eta \sim N(0, \sigma^2)$. With this assumption, the joint probability
8 expression for trip fare and trip destination may be derived. Conditional on η the probability for a
9 trip, q to have fare, y_q and destination s can be expressed as follows:

10

$$P(y_q, s_q | \eta) = P(y_q) \times P(s_q) \quad (5)$$

11

12 The complete set of parameters to be estimated in the model system of Equation (5) are
13 α, β and γ and standard error term, σ . Let, Ω represents a vector that includes all the standard
14 error parameters to be estimated. Given this assumption, the joint likelihood for trip fare and trip
15 destination is provided as follows:

16

$$L_q | \Omega = \prod_{s=1}^S [P(y_q, s_q | \eta)]^{d_{qs}} \quad (6)$$

17

18 where d_{qs} is a dummy variable taking a value of 1 if a user making the trip, q chooses the
19 destination, s and 0 otherwise. Finally, the unconditional likelihood function may be computed for
20 a trip, q as follows:

21

$$L_q = \int_{\Omega} (L_q | \Omega) d\Omega \quad (7)$$

22

23 Now, we can express the log-likelihood function of the final joint model as follows:

24

$$LL = \sum_{q=1}^Q \ln L_q \quad (8)$$

25

26 The log-likelihood function in Equation (8) involves the evaluation of a multi-dimensional
27 integral of size equal to the number of rows in Ω . We apply Quasi-Monte Carlo simulation
28 techniques based on the scrambled Halton sequence to approximate this integral in the likelihood
29 function and maximize the logarithm of the resulting simulated likelihood function (See Bhat, (54);
30 Yasmin and Eluru, (55) for more details).

31

32 **MODEL DEVELOPMENT**

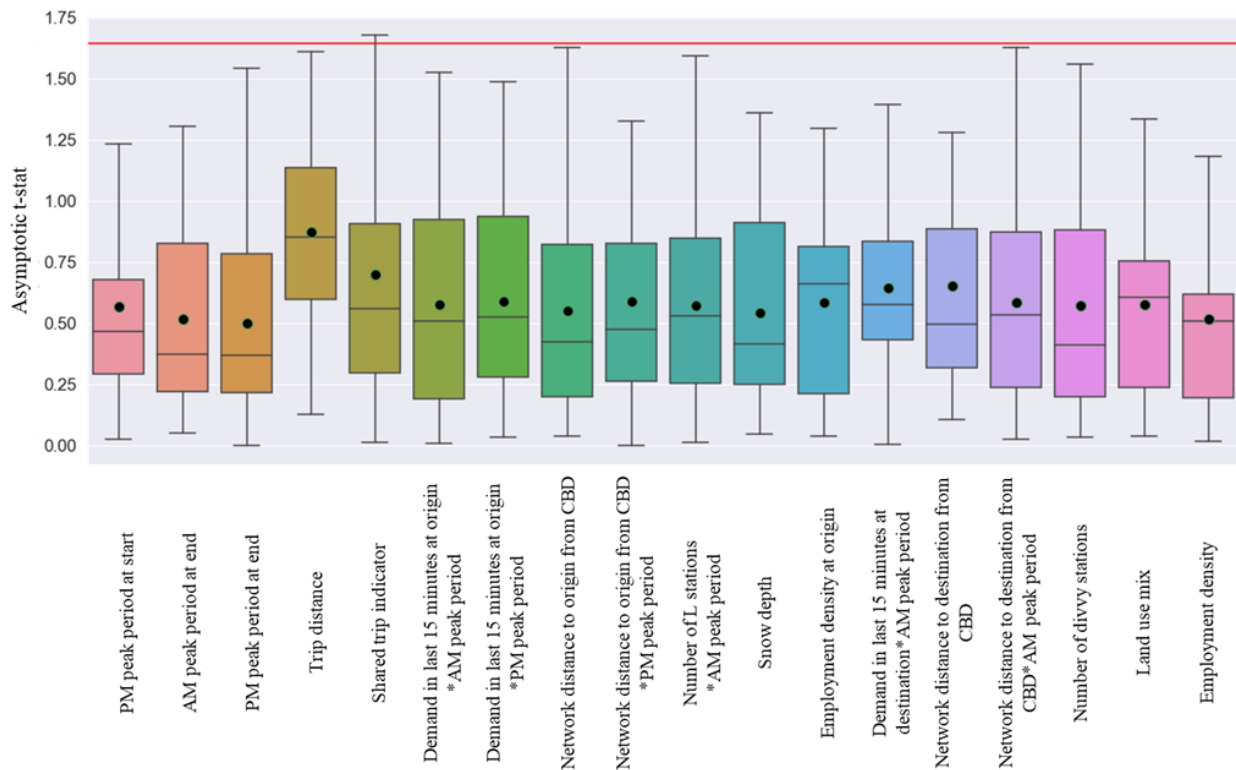
33 As described earlier, we estimate the model components employing a randomly chosen dataset of
34 10,000 records for computational efficiency and avoiding overfitting. Given the possibility that
35 the random sample might not represent the population, we draw 25 samples of 10,000 and examine
36 the role of randomness in the parameter stability across the samples for linear regression and

1 multinomial logit models. To examine parameters stability, we employ the following revised Wald
 2 test statistic approach across 25 samples:

$$3 \text{ Parameter test statistic} = \textit{abs} \left[\frac{(\textit{sample parameter} - \textit{population benchmark})}{\sqrt{SE_{\textit{sample}}^2 + SE_{\textit{population}}^2}} \right]$$

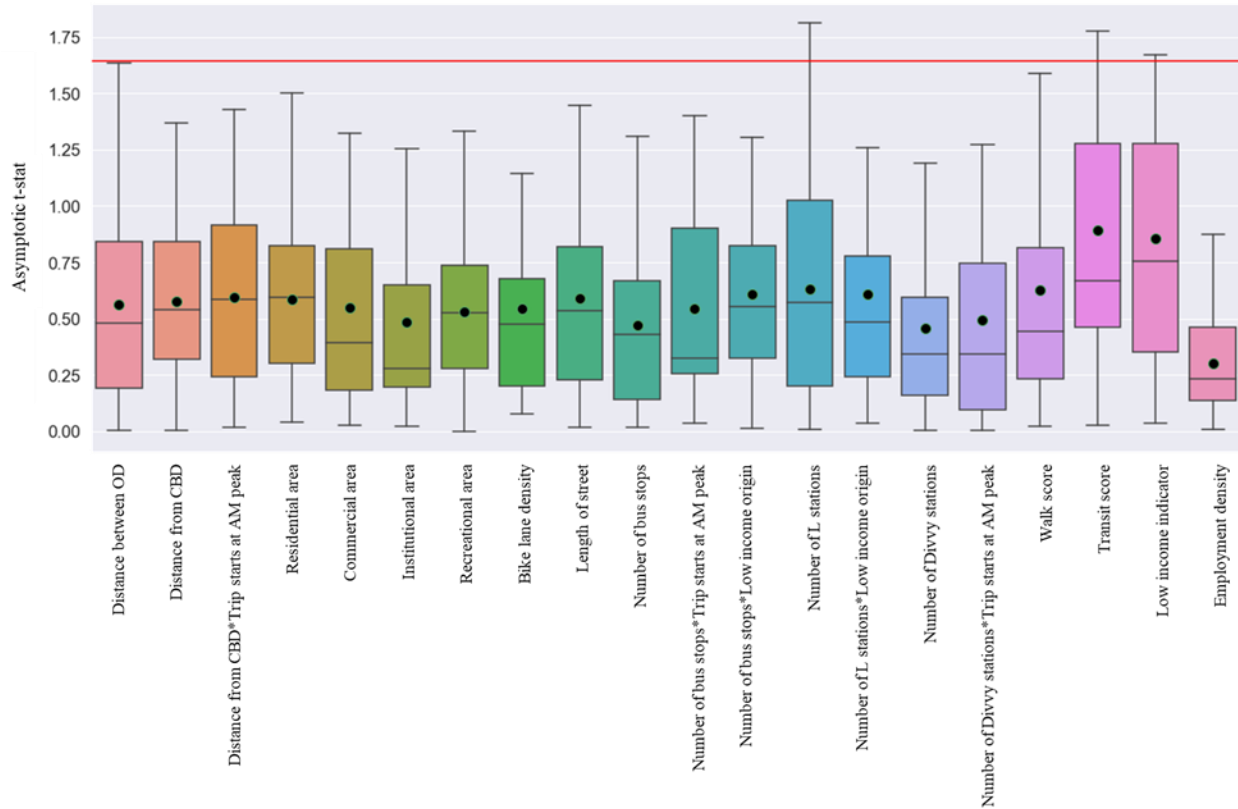
4
 5 The population benchmark is computed as the average value of the parameter across the
 6 25 samples. If any parameter for a sample is significantly different from the population benchmark,
 7 the Wald statistics will be larger than the 90% t-statistic value of 1.65. **Figure 2** and **Figure 3**
 8 illustrates the range of revised Wald test statistic for 25 samples for trip fare model and destination
 9 choice model respectively in a box plot. It is evident from **Figure 2** and **Figure 3** that means of
 10 the revised Wald test statistic of all the exogenous variables (and majority of the realizations) are
 11 well within 90% test statistic. To be precise, in case of the trip fare model (destination choice
 12 model), only two (six) test statistic values among 450 (500) values generated were found to be
 13 greater than 90% test statistic. Therefore, we can conclude that the parameters estimated across
 14 the random samples are stable and there is no significant difference in parameters estimated across
 15 samples.

16 After establishing that the random sample based models are stable, we estimate a joint LR-
 17 MNL model which accounts for common unobserved heterogeneity between trip fare and trip
 18 destination for one sample.



19

20 **Figure 2 Asymptotic t-statistic for the parameters estimated of trip fare model**



1
2 **Figure 3 Asymptotic t-statistic for the parameters estimated of destination choice model**

3
4 **MODEL ESTIMATION RESULTS**

5 The model performance of the proposed joint model is compared to the independent fare and
6 destination models using log-likelihood (LL) and Bayesian Information Criterion (BIC) measures.
7 The LL (BIC) values of the independent LR and MNL model are -22857.920 (46075.043). For the
8 joint LR-MNL model system, LL (BIC) values were found to be -22717.000 (45793.203). Hence,
9 the joint model system clearly outperforms the independent model system. For the sake of brevity,
10 the results from the Joint LR-MNL model estimation results are discussed (see **Table 2**). The
11 discussion is organized by variable group.

12
13 **Trip Fare Model**

14
15 *Trip Attributes*

16 In **Table 2** several trip attributes are found to have significant impact on TNC fare. Trip distance,
17 as expected, has a positive impact on trip fare. Controlling for everything else, longer trips have
18 higher fares. If the trip starts and ends in PM peak period, fare is likely to increase for the
19 corresponding trip. Similarly, when a trip ends in AM peak period, an increase in fare is observed.
20 The results are along expected lines and suggest that during peak periods a higher fare is levied.
21 Finally, we also find that shared trips are likely to have a lower fare as expected.

22
23 **TABLE 2 Joint LR-MNL Model Result**

| Variable | Estimate | t-stat |
|---|----------|---------|
| Trip Fare Model | | |
| Constant | 1.638 | 86.286 |
| Trip Attributes | | |
| Trip distance | 0.104 | 77.797 |
| Trip start time (Base: Other periods) | | |
| PM peak period | 0.096 | 3.546 |
| Trip end time (Base: Other periods) | | |
| AM peak period | 0.045 | 2.276 |
| PM peak period | 0.045 | 2.326 |
| Shared trip indicator (Base: No) | | |
| Yes | -0.371 | -37.494 |
| Origin Attributes | | |
| Demand in last 15 minutes at origin*AM peak period | -0.019 | -2.683 |
| Demand in last 15 minutes at origin*PM peak period | 0.021 | 2.683 |
| Network distance to origin from CBD | 0.010 | 5.271 |
| Network distance to origin from CBD*PM peak period | -0.011 | -2.713 |
| Number of L stations*AM peak period | -0.008 | -1.643 |
| Employment density at origin | -0.012 | -2.675 |
| Snow depth | -0.010 | -2.838 |
| Destination Attributes | | |
| Demand in last 15 minutes at destination*AM peak period | 0.032 | 5.157 |
| Network distance to destination from CBD | -0.007 | -3.836 |
| Network distance to destination from CBD*AM peak period | -0.005 | -2.335 |
| Number of Divvy stations | -0.004 | -4.535 |
| Land use mix | 0.183 | 1.959 |
| Employment density | 0.014 | 3.096 |
| Scale | 0.229 | 31.564 |
| Destination Choice Model | | |
| Land Use and Built Environment Attributes | | |
| Network distance between O-D | -1.092 | -57.340 |
| Distance from CBD | -0.539 | -16.967 |
| Distance from CBD*Trip starts at AM peak | -0.254 | -4.705 |
| Residential area | -0.927 | -14.191 |
| Commercial area | 0.396 | 7.397 |
| Institutional area | -0.168 | -2.939 |
| Recreational area | 0.306 | 7.262 |
| Transportation Infrastructure Attributes | | |
| Bike lane density | 0.160 | 4.685 |
| Street Length | 0.073 | 25.453 |
| Number of bus stops | 0.007 | 3.064 |

| Variable | Estimate | t-stat |
|---|----------|---------|
| Number of bus stops*Trip starts at AM peak | 0.020 | 5.408 |
| Number of bus stops*Low income origin | 0.017 | 4.692 |
| Number of L stations | -0.117 | -7.473 |
| Number of L stations*Low income origin | -0.120 | -3.726 |
| Number of Divvy stations | 0.037 | 4.465 |
| Number of Divvy stations*Trip starts at AM peak | 0.039 | 2.524 |
| Walk score | 0.004 | 4.158 |
| Transit score | 0.089 | 3.597 |
| Demographic Attributes | | |
| Low income indicator (Base: Median income over 15 th percentile) | | |
| Yes | -0.926 | -24.680 |
| Employment density | 0.061 | 2.242 |
| Unobserved heterogeneity | | |
| Constant in LR and Distance between O-D in MNL | 0.266 | 36.494 |
| Constant in LR and Street Length in MNL | 0.039 | 7.148 |

1

2 *Origin Attributes*

3 In our analysis, we wanted to consider the influence of demand in preceding time intervals on trip
4 fare. For this purpose, origin demand in the last 15 minutes in AM and PM peak periods was
5 considered in the model. The model estimates offer interesting results. In the AM peak period,
6 higher demand has a negative coefficient. While this might appear counter-intuitive on first glance,
7 the reader will recognize that the demand variable interacts with the AM peak main effect thus,
8 the net effect is still likely to be positive. For PM peak period, the impact on fare is more
9 pronounced clearly highlighting that higher demand at the origin contributes to a higher fare.

10 From **Table 2**, it is evident that trip fare is likely to increase as distance between origin of
11 the trip and CBD increases. The result represents the supply side challenge (or rerouting costs) for
12 drivers to pick up riders away from CBD (see (56) for similar findings). The negative coefficient
13 for interaction of distance variable and PM peak period indicates that during PM peak the impact
14 of distance from CBD is moderated potentially due to increase expected supply for TNC. Chicago
15 L, a rapid transit system, operates inside the city of Chicago. The number of L stations in the AM
16 peak period has negative impact on TNC fare highlighting potential competition between Chicago
17 L and TNC (57). Interestingly, higher employment density at the origin is negatively associated
18 with TNC fare potentially reflecting the presence of infrastructure for non-motorized modes and
19 improved land use (58, 59). The results indicate that in adverse weather conditions such as higher
20 level of snow depth, TNC fares are likely to be lower possibly due to supply demand imbalance
21 (60, 61).

22

23 *Destination Attributes*

24 The demand in the last minutes at the destination also offers interesting results. We find that
25 interaction of destination demand with AM peak is positive indicating that higher fares are likely
26 to destinations with higher demand in AM peak (similar findings in 11). As the distance of the
27 destination census tract increases from CBD, TNC fare is likely to be lower. The result is expected
28 because with all else same, travel away from CBD is typically faster and thus trip fare is expected

1 to be lower. The effect is more pronounced in the AM peak period as congestion is likely to be
2 lower away from CBD during AM peak.

3 Chicago bike sharing system (Divvy) and TNC appear to have competitive relationship as
4 highlighted by the negative coefficient on the number of divvy stations (see (30) for evidence of
5 how individuals use divvy system to make commuting trips in CBD). The results also indicate that
6 destination with diverse land use is likely to have higher fares. TNC travel in these locations will
7 be slower and hence require longer travel time resulting in higher fares. Finally, destinations with
8 higher employment density will contribute to higher TNC fare as expected.

9

10 **Destination Choice Model**

11 *Land use and Built Environment Attributes*

12 Several land use and built environment variables offer significant and expected results. As the
13 distance between origin and destination and distance of the destination from CBD increases, the
14 likelihood of the alternative being selected reduces. The impact of distance to CBD is significantly
15 higher in the AM peak period as users are unlikely to travel away from the CBD in the AM peak.
16 The various built-up areas also offer expected results. Census tracts with residential and
17 institutional areas are less likely to be destination. On the other hand, census tracts with higher
18 areas of commercial and recreational areas have a higher likelihood of being chosen (see (32, 63)
19 for similar results).

20 *Transportation Infrastructure Attributes*

21 The results for transportation infrastructure attributes offer multiple significant and nuanced
22 relationships with destination preferences. Destination attributes that represent non-motorized and
23 transit infrastructure such as bike lanes, bus stops, divvy stations, walk score and transit offer
24 positive association with destination choice. Several earlier studies have documented these some
25 or all of these relationships (21, 22, 60, 64–66). For bus stops and divvy stations, the impact on
26 destination selection is even higher during the AM peak period. An exception to this is the
27 parameter for L stations. The result clearly highlights that in census tracts with L stations, TNC
28 users are less likely to choose these destinations. The income of origin census tract also offers a
29 conflicting interaction with bus stops and L stations. The users starting their travel from low-
30 income census tracts have higher affinity to travel to destinations with higher number of bus stops.
31 However, the result is exactly opposite in the context of L stations. The variation might be
32 reflecting the different neighborhood characteristics of census tracts with higher number of buses
33 vis-à-vis census tracts with higher number of L stations (62, 67, 68).

34

35 *Demographic Attributes*

36 Census tracts with lower income are less likely to be chosen as TNC destinations. The result
37 indicates to income inequity in the adoption of TNC for mobility needs in Chicago and other urban
38 regions (see similar findings in (23, 69, 70). As expected, on weekdays, a census tract with higher
39 employment density is likely to attract more TNC trips (31, 69).

40

41 **Unobserved Heterogeneity**

42 The proposed LR-MNL joint model system accommodates for common unobserved heterogeneity
43 between trip fare and destination choices. Several unobserved factors were tested in the joint
44 model. The variables that offered significant unobserved correlation are reported in the last row
45 panel of **Table 2**. The two parameters represent interaction of a constant in fare model with origin
46 -destination distance and street length. These significant correlations reinforce our hypothesis that

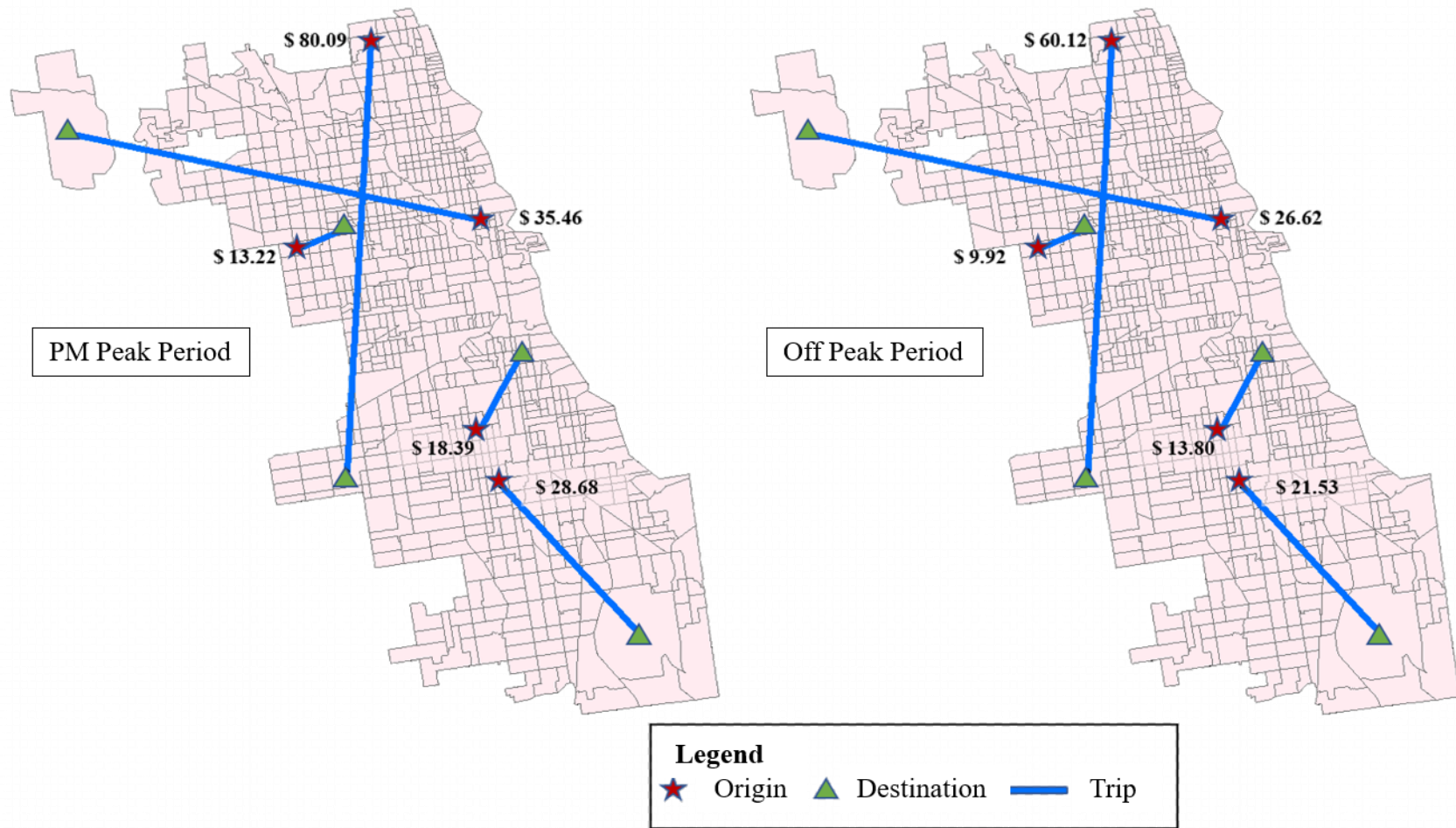
1 trip fare and destination choices are influenced by shared factors and incorporating such correlation
2 is important.

3

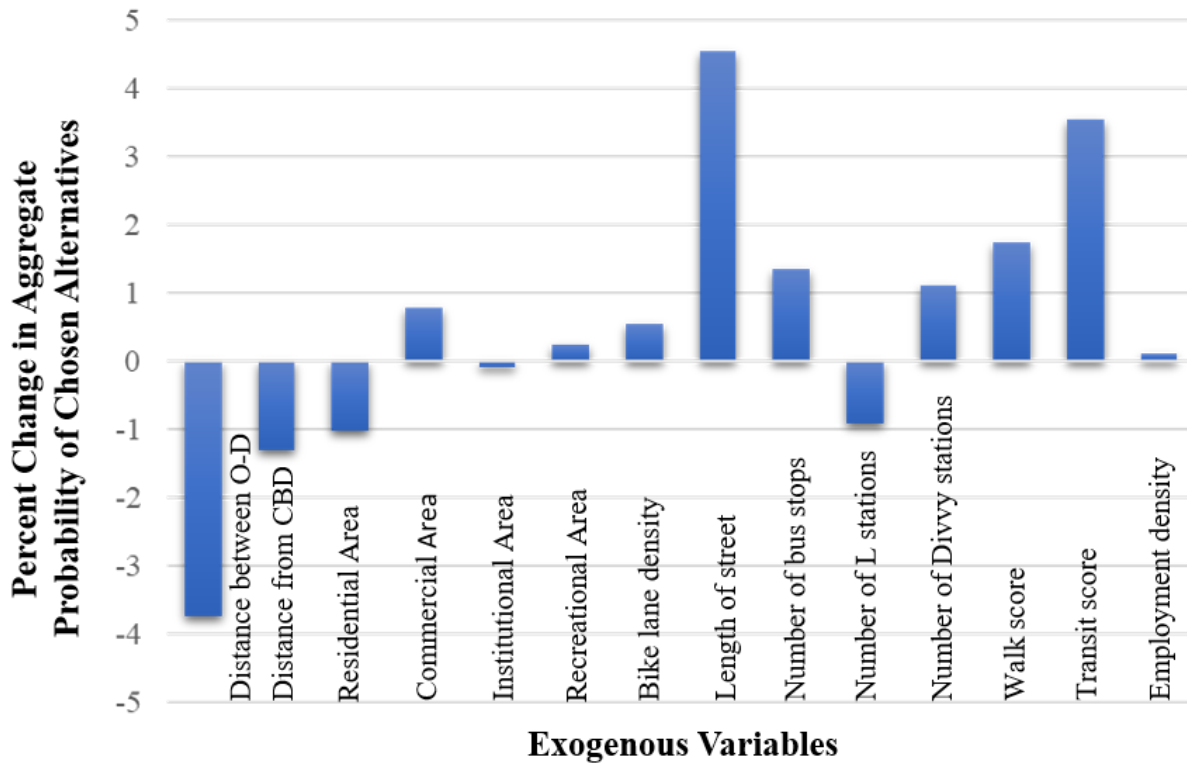
4 **PREDICTION AND ELASTICITY ANALYSIS**

5 To illustrate the applicability of the proposed model, we employ the model results for
6 understanding the influence of independent variables on fare and destination choice models. We
7 employ the model results for the fare model to generate trip cost predictions for five randomly
8 chosen trips in the PM peak and off-peak periods. These trips are plotted in **Figures 4 (a) and (b)**.
9 The prediction illustrates how the proposed model can be employed for generating trip fares across
10 the region. The trip fares presented in **Figure 4** illustrate the higher cost of TNC during PM peak
11 (relative to off-peak period). The procedure can be readily applied to generate travel cost schemes
12 for a mode choice model in the region with TNC alternative.

13 For the destination model, an elasticity analysis has been undertaken in an effort to capture
14 the changes in dependent variables (destination) in response to changes in independent variables.
15 **Figure 5** illustrate the percent change in fare and aggregate probability of the chosen destination
16 alternative respectively due to change in independent variables by 10%. The results summarized
17 in **Figure 5** offer interesting results. We notice that distance between origin destination, transit
18 score and street length variables exhibit the highest impact on destination preferences. We also
19 observe that walk score, divvy stations, bus stops, residential area and distance form CBD affect
20 destination preferences reasonably. In summary, the elasticity effect highlights how transportation
21 planners and TNC owners can examine trends influencing destination choice behavior.



1
2 **Figure 4 Trip fare prediction across (a) PM peak period; (b) Off peak period**



1
2 **Figure 5 Elasticity analysis**

3
4 **CONCLUSIONS**

5 Given the prevalence of Transportation Networking Companies (TNCs) across the world, there is
6 growing literature dedicated to TNC usage analysis. However, there is limited research on
7 comprehensively examining the influence of independent variables on TNC fare. In this study, we
8 postulate that TNC trip fare is closely linked to TNC trip destination and develop a joint
9 econometric model linking the two outcomes. A wide range of origin and destination specific land
10 use and built environment factors, transportation infrastructure attributes, and weather attributes
11 were found to be significant in the joint the model system. Based on log-likelihood (LL) and
12 Bayesian Information Criterion (BIC) measures, the model performance of the proposed joint
13 model is found to be superior compared to independent fare and destination models. The model
14 results were augmented with fare prediction exercise and destination model elasticity analysis. The
15 fare prediction exercise illustrated how the proposed model can be employed to generate TNC
16 travel costs for use in a mode choice model with TNC alternative. The destination elasticity
17 analysis highlighted the important factors affecting destination preferences.

18 The study is not without limitations. TNC trip data does not provide any user related
19 information. Access to sociodemographic, socioeconomic, and other relevant information can
20 significantly enhance the models developed in our analysis. Trip level TNC data employed in this
21 study provides trip origin and destination aggregated at the census tract level potentially to
22 preserve user and operator privacy. The aggregated destination information can result in large
23 differences in travel distances for short trips within the census tracts. The model developed can be
24 further refined in the presence of more disaggregate data. It is also important to recognize that
25 TNC trip fare can be influenced by business strategies of TNCs that are not readily declared

1 publicly. Understanding the effect of TNC business strategies might be an avenue for future
2 research.

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9

10 **AUTHOR CONTRIBUTION STATEMENT**

11 The authors confirm contribution to the paper as follows: study conception and design: Naveen
12 Eluru , Tanmoy Bhowmik, Sudipta Dey Tirtha, Dewan Ashraful Parvez; data collection: Dewan
13 Ashraful Parvez, Sudipta Dey Tirtha, Tanmoy Bhowmik, Naveen Eluru; model estimation: Dewan
14 Ashraful Parvez, Sudipta Dey Tirtha, Tanmoy Bhowmik, Naveen Eluru; analysis and
15 interpretation of results: Dewan Ashraful Parvez, Sudipta Dey Tirtha, Tanmoy Bhowmik, Naveen
16 Eluru; draft manuscript preparation: Dewan Ashraful Parvez, Sudipta Dey Tirtha, Tanmoy
17 Bhowmik, Naveen Eluru. All authors reviewed the results and approved the final version of the
18 manuscript.
19

20 **CONFLICT OF INTEREST STATEMENTS**

21 The authors do not have any conflicts of interest to declare.
22

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